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Distrito Federal, México

Available in: http://www.redalyc.org/articulo.oa?id=47415823009
Determining a Checkout Register Opening Policy to Maximize Profit In convenience Stores Chains

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ABSTRACT
A major concern for convenience store managers is lost sales due to balking. Convenience stores customers pay high margins expecting fast service. If waiting lines are too long for their tolerance level at their arrival, they balk and the sale is lost as a result. In order to reduce lost sales, the length of the waiting line is usually controlled by opening additional checkout registers when the number of customers standing in the line exceeds a specified number and maintaining them open until they are no longer needed. This paper presents an applied approach to model the probability that customers actually enter the store and define the optimal opening level \( \mathcal{N} \)-policy of the second checkout register based on several factors including the particular waiting line length tolerance level of usual customers, the average hourly arrival rate of customers to the store and the average gain per customer transaction. Several performance measures are computed. The total expected cost function per unit time is proposed to determine the optimal operating \( \mathcal{N} \)-policy at minimum cost. The model is applied in a real case of a convenience store chain.

Keywords: Convenience store, balking process, customer service model, queuing control, removable servers, N-policy.

1. Introduction
A major concern for service industries is lost sales due to balking or reneging of impatient customers. Balking is the phenomenon that occurs when a potential customer decides not to enter the waiting line because he/she considers it is too long. Reneging occurs when a customer decides to leave the waiting line after entering because the waiting time has exceeded his/her tolerance level.

Pazgal & Radas (2008) made an empirical study to compare balking and reneging behavior. They
found, that when the waiting line length is observable as is the case in convenience stores, very few participants renege and most participants balk when line is longer than some critical length determined by the customer.

Lost sales due to balking can be very high, for example, a 1% balking rate in a drive-through of a fast food restaurant can reduce the net income around $100 million dollars per year (Jones, 1999). Due to this, it is very important to develop efficient methods that allow estimating and reducing lost sales caused by balking.

A common approach in convenience stores is to have more checkout registers available to open when the waiting line is too long. This measure reduces lost sales but requires more personnel to operate the additional checkout registers, thus increasing labor cost. The problem is to find the checkout register opening policy that reduces the total cost: the cost of making the customer wait plus the cost of providing the service (Budnick, 1977), (Fitzsimmons, 1982), (Davis, 1991).

This decision could appear simple and there exists several models to address this problem in literature based on a small set of variables such as average arrival rate and predefined balking models, among others. However, in a practical case, the convenience store manager must consider many other elements in order to find the right policy. These elements include the particular waiting tolerance of the usual customers in the area, the behavior of the arrival rate during the day, the average labor cost, and the average profit per transaction in the store.

In this paper we address this problem from the perspective of the manager of a real convenience store chain with hundreds of stores. Given that an average increase of a single dollar per hour per store can represent a profit of millions of dollars for the company, the focus is on finding a practical, precise and effective approach to find the best policy to manage the checkout registers.

2. Related literature

Several authors model customer impatience introducing cost functions rather than balking. Yadin and Naor (1963) consider a cost function that includes the cost of waiting and the cost of switching the number of servers. They first introduced the concept of an $N$-policy which turns the server on when $(N \geq 1)$ or more customers are present and turns the server off when no more customers are waiting for service. Steady-state probabilities and some performance measures were derived. Gebhard (1967) considers two particular service-rate-switching policies, assuming a Poisson input and exponential service times. The first policy is a single level control and the second one, a bilevel hysteretic control. The bilevel hysteretic control is stated as follows: when the system size reaches a level of $N_2$ from below, use rate $\mu_2$, and when the system size drops to $N_1$, switch to $\mu_1$. He compares both policies for specific cost functions that include service and waiting line costs. Heyman (1968) analyzes an M/G/1 system with removable server under several discount cost over time. Sobel (1969) generalizes it to G/G/1. Several authors. Bell (1979), Boots and Tijms (1999), Feinberg and Kella (2002) and Wang (2003) present more queue control strategies based on similar cost functions.

Other authors model customer impatience including balking or reneging models. Naor (1969) and Hassin (1986) present a model with balking in which customers will join the waiting line only if its length is less than a fixed constant that depends on the system utilizations factor, the reward obtained upon completion of service and the customer waiting cost. Van Tilts (1979) assumes that the probability that a customer enters the waiting line when there are $n$ customers in it is $b_n = 1/(n + 1)$.

Gross and Harris (1985) analyze M/M/m queuing systems with balking assuming a general balking probability function $(1 - b_n)$ when there are $n$ customers in a single line. Some common models for staying probability $b_n$ are summarized. One model assumes a common balking level $K$ for all customers as assumed in M/M/1/K model; however, “rarely do all customers have the same discouragement limit all the time” (Gross, 1985). Other models take the form of monotonically decreasing functions on the waiting line length $n$ as $1/(n + 1), 1/(n^2 + 1)$ and $e^{-an}/\mu$. They also present the $(c_1, c_2, N_1, N_2)$ policy, where $N_2 < N_1$. When the length of the waiting line reaches $N_1$, a
shift from $c_1$ to $c_2$ servers is done. When the waiting line falls to level $N_2$, we shift back to $c_1$.

Bae, Kim & Lee (2001) model balking assuming that a customer enters only if the expected waiting time is less than a fixed bound, common to all customers. Wang & Chang (2002) propose a staying probability that depends on system utilization and the number of customers waiting.

The aim of this article is to present a new model for balking probabilities that can be fitted to the particular impatience characteristics of the customers of a convenience store and can be used to model the waiting line behavior. It is used to estimate the expected loss profit per balking and determine the policy to open checkout registers that minimizes the total cost under the practical environment found in real convenience stores, including the changes in the demand rate during the day and the particular impatience model for its customers.

3. System Description

Consider a convenience store. Let \( N = \{ N_t; t \geq 0 \} \) be the customer arrival process. Customers usually arrive one at the time. When more than one customer arrives at the same time, we assume that only one pays and we count them as a single arrival. There are so many different customers that we can assume that arrival times are independent; therefore, we can assume that the mapping \( t \rightarrow N_t(w) \) has jumps of unit magnitude for almost all \( w \), and that for any \( t, s \geq 0, N_{t+s} - N_t \) is independent of \( \{N_u; u \leq t\} \). Due to the above, it can be assumed that the arrival process is a Non-Stationary Poisson Process (Cinlar, 1975).

The customer arrival rate changes during the day. If, for modeling purposes, we split the day in short enough time intervals, we can assume that the rate is constant within each interval. Without loss of generality, we will assume intervals of one hour. Let \( \lambda_t \) denote the arrival rate of customers during hour \( h \).

The service time of a customer in each checkout register is assumed to be exponential with media \( t_s \), then service rate is \( \mu = 1/t_s \). Also, we define the System Utilization assuming one checkout register opens as \( \rho = \lambda_t/\mu \). Each customer that enters the system (convenience store) observes the waiting line and based on its size, decides whether or not to enter. If the customer decides not to enter (balks), the sale is lost. The waiting line length at which a customer balks depends on his/her particular tolerance level to waiting. Let \( P(n) \) denote the staying probability, the probability that a customer does not balk and enters the convenience store when the number of customers in the waiting line is \( n \). The behavior of \( P(n) \) varies among stores depending on the tolerance level to waiting of the usual customers and the proximity of another convenience stores.

To reduce lost sales, the convenience store has two checkout registers and at least two employees. One of the checkout registers is permanently opened and operated by one employee with service rate \( \mu \). A second employee does product replenishment, cleaning and complementary tasks during low demand periods and opens and operates the second checkout counter when the number of customers in waiting line \( n \), reaches certain level which we will call the opening level \( \mathcal{N} \). The second checkout register remains open until it is no longer needed, when \( n = 1 \).

If the opening level \( \mathcal{N} \) for the second checkout register is low, the customer service level will be higher but the presence of the operator at the second checkout counter will be required more frequently, reducing the time available for complementary tasks. This may eventually require more employees, increasing the cost. On the other hand, if the opening level \( \mathcal{N} \) for the second checkout register is high, the labor cost of this checkout register will be reduced but the service level will also be reduced. As a consequence, lost sales due to balking would increase.

The objective is to determine the optimal opening level \( \mathcal{N}^* \) for the second checkout register in order to minimize the total cost (lost sales plus labor cost).

4. Modeling balking behavior

A sampling study was performed to obtain empirical data about balking behavior. Models found in reviewing literature did not fit well to the observed data. Empirical values of \( P(n) \) could be used if there were enough information. However, the
sample size needs to be quite large in order to observe the monotone decreasing values that are expected.

A mathematical function that could be adjusted to fit the observed behavior was proposed and the parameters of the model were obtained to validate the goodness of fit.

Let \( P(\text{Staying}|n_{pck}) \) denote the probability that a customer enters the store when the average number of customers per checkout register is \( n_{pck} \).

The proposed model is

\[
P(\text{Staying}|n_{pck}) = P(n_{pck}) = \frac{1}{\left(\frac{n_{pck}}{N}\right)^{N+1}}
\]  

(1)

Where \( N \) and \( \alpha \) are parameters of the model that provide flexibility to fit the empirical data.

\( N \) is the indifference level, the waiting line length per checkout register at which the probability to enter or balk is equal (Observe that when \( N = n_{pck} \), the probability of staying is equal to 0.5).

\( \alpha \) is the model parameter that adjusts the rate of probability of declining close to \( N \).

Let \( P(k) \) be the observed staying probability when the waiting line length per checkout counter is \( k \) \( (n_{pck} = k) \) and \( o_k \) be the number of observations used to compute \( P(k) \). Let \( \theta = (\alpha, \beta) \) denote a set of model parameters, and \( \hat{P}(k, \theta) \) the probability predicted by the model. We define the Weighted Sum of Square Errors as

\[
WSSE(\theta) = \sum_{k=1}^{\infty} \left(P(k) - \hat{P}(k, \theta)\right)^2 \cdot o_k
\]  

(2)

Where the \( \theta^* = (\alpha, \beta) \) that minimizes \( WSSE \) can be obtained using a numerical fitting algorithm.

To validate the adequacy of the model, a real convenience store was sampled. At the time of each arrival, the decision of the customer (staying or balking), the number of checkout registers open, and the number of customers waiting in line was recorded. The parameters that minimize \( WSSE \) were obtained: \( N = 9.46 \) y \( \alpha = 4.78 \). Figure 2 compares the observed and predicted probability.

The number of observations for the highest values of the waiting line length was small because during the observed period the waiting lines rarely reached those values. This explains the high variability in the last values.
5. Modeling Waiting line Dynamics

An approach based on Markov process analysis was used to model the waiting line dynamics. We define the following random variables as follows:

\[ Y_t = \text{occupancy (number of customers waiting in line at checkout register one)} \]
\[ n_{cr_t} = \text{number of checkout registers opened at instant } t, \forall t \in N \]

We suppose that the arrival process can be approximated as a Poisson process with arrival rates \( \lambda_t \) constant per interval. We define the stochastic process as \( \mathcal{Y} = \{Y_t, n_{cr_t}\}, \forall t \in R \}. Given that the arrival stream is assumed Poisson and service times are assumed exponentially distributed, the \( \mathcal{Y} \) process is a Markov Process.

\[ P_{ij} = P\{\text{Customer enters the system}|Y_t = i, n_{cr_t} = j = 1j/N+x+1 \} \tag{3} \]

Figure 2. Observed and Predicted Staying Probability

Figure 3. Observed and Predicted Staying Probability.
Figure 3 presents the transition diagram of the \( \mathcal{Y} \) process during hour \( h \). The circles represent the states \((Y_t, n_{cr_t})\) among which the process can evolve and the values in the arches represent the rates at which the process "jumps" between one state and another. Observe that the effective arrival rate when the state of the system is \((i, j)\) is \( \lambda_{\text{eff}} = P_i \lambda_h \) and \( \lambda_h (1 - P_i) \) represents the dropout rate. We define \( n_{\text{max}} \) as the maximum amount of customers that the system can handle and \( \mu \) as the service rate of a single checkout register (customers per hour).

This system can be analyzed using standard Markov Processes results. \( G \) represents the Generator Matrix of the \( \mathcal{Y} \) Process. Define \( \pi^h \) as the steady-state probability distribution of the \( \mathcal{Y} \) process at hour \( h \). Then \( \pi^h \) can be obtained solving the following system:

\[
\pi^h G = 0
\]

\[
\sum_{k \in E} \pi^h(k) = 1
\]

6. Performance Measures

The steady-state probability distribution of the \( \mathcal{Y} \) process at hour \( h \) defined as \( \pi^h \) is used to compute several performance measures: Expected value of lost sales, cashier cost and optimal opening level function \( N^*(\rho) \).

When the system is in state \((i, j)\), the probability that a customer enters the system is \( P(i, j) \). Then, the probability of losing a customer is \( 1 - P(i, j) \). Let \( P^h_{ic} \) represent the probability of losing a customer arriving during hour \( h \). Then,

\[
P^h_{ic} = \sum_{(i,j) \in E} \pi^h(i,j) (1 - P(i,j))
\]

Given that \( U_G \) is the average unitary gain per customer, the expected profit lost per customer balking during hour \( h \) is

\[
E[LP_h] = E[\text{Lost Profit in hour } h] = P^h_{ic} \lambda_h U_G
\]

Let \( N^h_{cr} \) be the expected number of checkout register open during hour \( h \),

\[
N^h_{cr} = 1 \sum_{i} P_i (1 + 2 \sum_{vi} \pi^h(i, 2))
\]

Let \( C_i \) be the labor cost per employee working hour. The cost of operating the checkout registers during hour \( h \) is

\[
E[CC_h] = E[\text{Cashiers Cost in hour } h] = C_i N^h_{cr}
\]

The total cost to run the store during hour \( h \) would be the sum of the cashiers cost plus the lost profit, then,

\[
E[T_C_h] = E[\text{Total Cost per Hour } h] = E[LP_h] + E[CC_h]
\]

Both costs depend on the opening level \( N \) of the second checkout register but the effects are in opposite directions. Increasing \( N \) increases the average lost profit but decreases the required number of cashiers. Decreasing \( N \) decreases the average lost profit but increases the average number of cashiers required.

It is necessary to determine the optimal opening level \( N^* \) that minimizes the expected value of the total cost. In order to obtain this value for hour \( h \), \( E[T_C_h] \) is obtained for several values of \( N \) and the value that minimizes the total cost is optimal value \( N^* \).

In order to avoid computing the optimal \( N^* \) value for each particular case, it is better to compute function \( N^*(\rho) \) that maps utilization \( \rho \) to the optimal \( N^* \) value. Then, in each particular case (one specific hour in one specific store in which the staying probability model applies) all that is necessary is to compute utilization \( \rho_h = \lambda_h / \mu \) and to identify the optimal \( N^* \) value from function \( N^*(\rho) \). A typical plot of this function is shown in figure 9. It is based on data from the case study presented in the next section.

In practice, choosing different opening levels for each hour could be complicated to manage. In this
case, an alternative approach is to choose a single $N$ level for the day (selecting the $N$ value that minimizes the total cost per day) or use one level per shift during the day and a different $N$ level at night according to the arrival rates in each period.

7. Results of the practical case study

The proposed approach was implemented in a real convenience store in Mexico using a computational code specially designed for this purpose. The arrival rate changes with the day and hour of the day. It affects the utilization per hour $\rho_h = \lambda_h / \mu$. Now, we will analyze the behavior of the performance indexes for several $\rho$ values. We assume an average sale of 2.50 dollars and a unitary gain of 1.00 dollars per sale. We suppose also that labor cost $c_l$ is 2.50 dollars per hour.

Figure 4 presents the total cost per hour per store (sum of labor cost plus sale loss) for different values of the opening level $N$ of the second checkout register and different utilization values.

Observe that the curve of minimum cost depends on the utilization. The optimal value is shown in figure 5.

Figure 5. Optimal opening level for second checkout register.
For utilization levels close to 2 or higher, we suggest analyzing the incorporation of a third checkout register in the model. It is not considered at this point. These results are valid under the cost relationship established and for the staying probability considered.

Figure 6 matches optimal opening level $N^*$ (shown on the right side of the plot) with the transactions per hour during the week in a real store.

A detailed policy could be to use the opening level indicated in Figure 6 for each hour. In practice, the store manager could decide to use a simpler policy that uses the same value for longer periods of time, for example:

- During the first two shifts, between 6:00 a.m. and 10:00 p.m., use $N = 4$, except from 6:00 to 8:00, from Monday to Friday, where $N = 2$ should be used due to the recurring peak. It is suggested to perform cleaning and replenishment activities before this hour in order to have available both checkout registers (cashiers) available at peak hours.
- Between 10:00 p.m. and 6:00 a.m., use $N = 6$.

8. Conclusions and managerial implications

The approach proposed in this paper has been developed from the perspective of the manager of a convenience store chain with hundreds of stores. Given that a single dollar in lost sales per hour can represent millions of dollars for the company, the focus has been more in precision and practical applicability rather than on simplicity.

The mathematical function proposed to model the probability that a customer actually enters the system approximates the observed staying probabilities in real cases in best way than other models found in literature.

The proposed approach to predict the dynamic behavior of the system allows estimating key performance measures of the system and determining the optimal opening value of the second checkout register for a convenience store for a specific hour.

The model is also used to model the dynamic behavior of the system for different values of the opening level and different levels of utilization. It allows estimating the lost sales, the number of checkout registers required, and the total cost for different combinations of $N$ values and levels of utilization $\rho$.

The model is used to decide the best practical policy to open a second checkout register throughout the entire week. The model can also be easily extended to more checkout registers. Finally, a practical case was presented to demonstrate the applicability of the model.
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Acknowledgments

This research and divulgation has been supported by ITESM Campus Monterrey Research Fund CAT128.
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